

Derivation of Markov Process from Path Entropy Maximization

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OUTLINE



- ❖ Maximum Entropy Principle
- ❖ Path Entropy Maximization
- ❖ Derivation of Markov process

Maximum Entropy Principle

- ❖ E.T. Jaynes (Phys. Rev. 106, 620(1957), Phys. Rev. 108, 171(1957)): Statistical Mechanics as a logical inference
- ❖ Maximize Gibbs–Shannon entropy $-\sum p_i \log p_i$ under given constraints : Most unbiased estimate

Boltzmann distribution



❖ Constraint: mean energy $\sum p_i E_i = \epsilon$

❖ Normalization: $\sum p_i = 1$

❖ In the absence of other information (equilibrium), the most unbiased estimate of the probability distribution is obtained by maximizing

$$-\sum_i p_i \log p_i + \beta \left(\sum_i p_i E_i - \epsilon \right) + \nu \left(\sum_i p_i - 1 \right)$$

Boltzmann distribution



$$- \sum_i p_i \log p_i + \beta \left(\sum_i p_i E_i - \epsilon \right) + \nu \left(\sum_i p_i - 1 \right)$$

$$\delta p_i : -\ln p_i - 1 + \beta E_i + \nu = 0$$

$$\delta \beta : \sum_j p_j E_j - \epsilon = 0$$

$$\delta \nu : \sum_j p_j - 1 = 0$$

Boltzmann distribution



$$p_i = \frac{e^{-\beta E_i}}{\sum_j e^{-\beta E_j}}$$

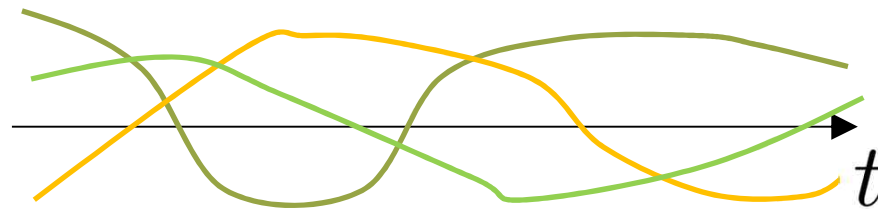
With β determined by

$$\frac{\sum_k E_k e^{-\beta E_k}}{\sum_j e^{-\beta E_j}} = \epsilon$$

Path entropy maximization



- ❖ Dynamical system: Obtain probability distribution $P(C)$ for path C



- ❖ Maximize the path entropy $-\sum_C P(C) \ln P(C)$ under appropriate constraints

$$\sum_C P(C) A^{(\alpha)}(C) = A_0^{(\alpha)} \quad \text{and} \quad \sum_C P(C) = 1$$

Path entropy maximization



- ❖ Jaynes (“Macroscopic prediction”, in Complex Systems Operational Approaches in Neurobiology, Physics, and Computers, edited by H. Haken (Springer–Verlag, Berlin, 1985)):
“Maximum Caliber principle”
- ❖ Filyukov and Karpov (J. Eng. Phys. 13, 624, 1967; 13, 798, 1967)
- ❖ Filyukov (Eng. Phys. Thermophys. 14, 814, 1968)

Discrete time (Filyukov et al.)



❖ $P(C) = p(i_0, i_1, \dots, i_T)$

❖ Path entropy: $H(T) = - \sum_{i_0, i_1, \dots, i_T} p_{i_0 i_1 \dots i_T} \log p_{i_0 i_1 \dots i_T}$

❖ Stationary Markov process

$$p_C = p_{i_0} p_{i_0 \rightarrow i_1} p_{i_1 \rightarrow i_2} \dots p_{i_{T-1} \rightarrow i_T}$$

$$H(T) = - \sum_i p_i \log p_i - T \sum_{i,j} p_i p_{i \rightarrow j} \log p_{i \rightarrow j}$$

❖ However, Markovian property itself can be derived from the property of the constraints!

Markov processes: Definitions



❖ n-point marginal probability

$$p(a_1, \dots, a_n; t) \equiv \sum_{i_0, i_1, \dots, i_{t-n}, j_1, j_2, \dots, j_{T-t}} p(i_0, i_1, \dots, i_{t-n}, a_1, \dots, a_n, j_1, j_2, \dots, j_{T-t})$$

❖ Transition probability

$$p(i_0, \dots, i_{t-1} \rightarrow i_t) \equiv \frac{p(i_0, \dots, i_t)}{p(i_0, \dots, i_{t-1})}$$

❖ n-th order Markov process

$$p(i_0, \dots, i_{t-1} \rightarrow i_t) = p(i_{t-n}, i_{t-n+1} \dots i_{t-1} \rightarrow i_t; t) \equiv \frac{p(i_{t-n}, \dots, i_t; t)}{p(i_{t-n}, \dots, i_{t-1}; t)}$$

→ Transition probability depends only on previous n steps of history

Derivation of Markov processes



- ❖ n-th order Markov process follows if only up to (n+1)-point function is constrained
- ❖ General constraints:

$$\sum_C P(C) A^{(\alpha)}(C) = \sum_{\{i_0, i_1, \dots, i_T\}} p(i_0, i_1, \dots, i_T) A^{(\alpha)}(i_0, i_1, \dots, i_T) = A_0^{(\alpha)}$$

- ❖ One-point constraints:

$$A^{(\alpha)}(i_0, i_1, \dots, i_T) = \sum_{t=0}^T \varepsilon_{i_t}^{(\alpha)}$$

$$\begin{aligned} F_0^{(\alpha)} &\equiv \sum_{\{i_0, i_1, \dots, i_T\}} \left(\sum_{t=0}^T \varepsilon_{i_t}^{(\alpha)} \right) p(i_0, i_1, \dots, i_T) - (T+1) E_0^{(\alpha)} \\ &= \sum_{t=0}^T \sum_{i_t} \varepsilon_{i_t}^{(\alpha)} p(i_t; t) - (T+1) E_0^{(\alpha)} = 0 \quad (\alpha = 1, \dots, N_1) \end{aligned}$$

Derivation of Markov processes



❖ Two-point constraints:
$$A^{(\alpha)}(i_0, i_1, \dots, i_T) = \sum_{t=0}^{T-1} J_{i_t i_{t+1}}^{(\alpha)}$$

$$\begin{aligned} F_1^{(\alpha)} &\equiv \sum_{\{i_0, i_1, \dots, i_T\}} \left(\sum_{t=0}^{T-1} J_{i_t i_{t+1}}^{(\alpha)} \right) p(i_0, i_1, \dots, i_T) - T J_0^{(\alpha)} \\ &= \sum_{t=0}^{T-1} \sum_{i_t i_{t+1}} J_{i_t i_{t+1}}^{(\alpha)} p(i_t, i_{t+1}; t) - T J_0^{(\alpha)} = 0. \quad (\alpha = 1, \dots, N_2) \end{aligned}$$

Derivation of Markov processes



❖ Take the variation of

$$\begin{aligned}
 & - \sum_{\{i_0, i_1, \dots, i_T\}} p(i_0, i_1, \dots, i_T) \log p(i_0, i_1, \dots, i_T) - \sum_{\alpha=1}^{N_1} \beta_{\alpha} \left(\sum_{t=0}^T \sum_{i_t} \varepsilon_{i_t}^{(\alpha)} p(i_t; t) - (T+1) E_0^{(\alpha)} \right) \\
 & + \sum_{\gamma=1}^{N_2} \nu_{\gamma} \left(\sum_{t=0}^{T-1} \sum_{i_t i_{t+1}} J_{i_t i_{t+1}}^{(\gamma)} p(i_t, i_{t+1}; t+1) - T J_0^{(\gamma)} \right) + (\rho + 1) \left(\sum_{\{i_0, i_1, \dots, i_T\}} p(i_0, i_1, \dots, i_T) - 1 \right)
 \end{aligned}$$

$$\delta p : -\log p(i_0, i_1, \dots, i_T) - \sum_{\alpha} \beta_{\alpha} \sum_{t=0}^T \varepsilon_{i_t}^{(\alpha)} + \sum_{\gamma} \nu_{\gamma} \sum_{t=0}^{T-1} J_{i_t i_{t+1}}^{(\gamma)} + \rho = 0$$

$$p(i_0, i_1, \dots, i_T) = \exp \left(\rho - \sum_{\alpha} \beta_{\alpha} \sum_{t=0}^T \varepsilon_{i_t}^{(\alpha)} + \sum_{\gamma} \nu_{\gamma} \sum_{t=0}^{T-1} J_{i_t i_{t+1}}^{(\gamma)} \right)$$

Derivation of Markov processes



$$p(i_0, i_1, \dots, i_T) = \frac{v(i_0)G(i_0, i_1)G(i_1, i_2) \cdots G(i_{T-1}, i_T)v(i_T)}{\mathbf{v}^\dagger \mathbf{G}^T \mathbf{v}}$$

$$v(i) \equiv \exp\left(-\sum_{\alpha} \beta_{\alpha} \varepsilon_i^{(\alpha)} / 2\right)$$

$$G(i, j) \equiv \exp\left(-\sum_{\alpha} \beta_{\alpha} \varepsilon_i^{(\alpha)} / 2 + \sum_{\gamma} \nu_{\gamma} J_{ij}^{(\gamma)} - \sum_{\alpha} \beta_{\alpha} \varepsilon_j^{(\alpha)} / 2\right)$$

Derivation of Markov processes



$$\begin{aligned} p(a_1, \dots, a_m; t) &= \sum_{i_0, \dots, i_{t-m}, i_{t+1}, \dots, i_T} p(i_0, i_1, \dots, i_{t-m}, a_1, \dots, a_m, i_{t+1}, \dots, i_T) \\ &= \frac{[\mathbf{v}^\dagger \mathbf{G}^{t-m+1}](a_1) G(a_1, a_2) G(a_2, a_3) \cdots G(a_{m-1}, a_m) [\mathbf{G}^{T-t} \mathbf{v}](a_m)}{\mathbf{v}^\dagger \mathbf{G}^T \mathbf{v}} \end{aligned}$$

$$\begin{aligned} p(a_1, \dots, a_m \rightarrow a_{m+1}; t) &= \frac{[\mathbf{v}^\dagger \mathbf{G}^{t-m}](a_1) G(a_1, a_2) \cdots G(a_m, a_{m+1}) [\mathbf{G}^{T-t} \mathbf{v}](a_{m+1})}{[\mathbf{v}^\dagger \mathbf{G}^{t-m}](a_1) G(a_1, a_2) \cdots G(a_{m-1}, a_m) [\mathbf{G}^{T-t+1} \mathbf{v}](a_m)} \\ &= \frac{G(a_m, a_{m+1}) [\mathbf{G}^{T-t} \mathbf{v}](a_{m+1})}{[\mathbf{G}^{T-t+1} \mathbf{v}](a_m)} = p(a_m \rightarrow a_{m+1}; t). \end{aligned}$$

Perron–Frobenius Theorem



(1) A positive matrix G has a positive real eigenvalue r , such that any other eigenvalue λ is strictly smaller than r in absolute value, $|\lambda| < r$.

(2) There is a left eigenvector $\mathbf{y}^\dagger = (y_1, \dots, y_N)$ for r with positive components. That is, $\mathbf{y}^\dagger \mathbf{G} = r\mathbf{y}^\dagger$ and $y_i > 0$ for all i . Similarly, there is a right eigenvector \mathbf{z} with positive components, such that $\mathbf{G}\mathbf{z} = r\mathbf{z}$ and $z_i > 0$ for all i .

(3) Left and right eigenvectors with eigenvalue r are non-degenerate.

$$(4) \lim_{T \rightarrow \infty} \frac{\mathbf{G}^T}{r^T} = \mathbf{z}\mathbf{y}^\dagger$$

Time homogeneity



$$T - t \rightarrow \infty$$

$$p(a \rightarrow b; t) = \frac{G(a, b)[\mathbf{G}^{T-t}\mathbf{v}](b)}{[\mathbf{G}^{T-t+1}\mathbf{v}](a)} \rightarrow \frac{G(a, b)z(b)}{rz(a)}$$

$$t \rightarrow \infty$$

$$p(a; t) = \frac{[\mathbf{v}^\dagger \mathbf{G}^t](a)z(a)}{r^t \mathbf{v}^\dagger \mathbf{z}} \rightarrow y(a)z(a)$$

→ Stationary Markov Process

Time homogenous master equation with an arbitrary initial distribution

- ❖ Initial condition is an additional information, which should also be implemented as a constraint

$$p(a; t = \tau) = \pi(a)$$

- ❖ Take variation of

$$\begin{aligned}
 & - \sum_{\{i_0, i_1, \dots, i_T\}} p(i_0, i_1, \dots, i_T) \log p(i_0, i_1, \dots, i_T) - \sum_{\alpha=1}^{N_1} \beta_{\alpha} \left(\sum_{t=0}^T \sum_{i_t} \varepsilon_{i_t}^{(\alpha)} p(i_t; t) - (T+1) E_0^{(\alpha)} \right) \\
 & + \sum_{\gamma=1}^{N_2} \nu_{\gamma} \left(\sum_{t=0}^{T-1} \sum_{i_t, i_{t+1}} J_{i_t i_{t+1}}^{(\gamma)} p(i_t, i_{t+1}; t+1) - T J_0^{(\gamma)} \right) + (\rho+1) \left(\sum_{\{i_0, i_1, \dots, i_T\}} p(i_0, i_1, \dots, i_T) - 1 \right) \\
 & + \sum_a \lambda(a) (p(a; \tau) - \pi(a))
 \end{aligned}$$

Time homogenous master equation

$$\begin{aligned}
 p(i_0, i_1, \dots, i_T) &= \exp(\rho + \lambda(i_\tau) - \beta \sum_{t=0}^T \varepsilon_{i_t} + \nu \sum_{t=0}^{T-1} J_{i_t i_{t+1}}) \\
 &= \exp(\rho + \lambda(i_\tau)) v(i_0) G(i_0, i_1) G(i_1, i_2) \cdots G(i_{T-1}, i_T) v(i_T) \\
 &= \frac{v(i_0) \pi(i_\tau) G(i_0, i_1) G(i_1, i_2) \cdots G(i_{T-1}, i_T) v(i_T)}{\sum_{j_0 \dots j_T} v(j_0) \pi(j_\tau) G(j_0, j_1) G(j_1, j_2) \cdots G(j_{T-1}, j_T) v(j_T)}
 \end{aligned}$$

$\tau < t$:

$$p(a_1, \dots, a_m \rightarrow a_{m+1}; t) = \frac{G(a_m, a_{m+1}) [\mathbf{G}^{T-t} \mathbf{v}](a_{m+1})}{[\mathbf{G}^{T-t+1} \mathbf{v}](a_m)}$$

$\tau \geq t$:

$$p(a_1, \dots, a_m \rightarrow a_{m+1}; t) = \frac{G(a_m, a_{m+1}) \sum_a [\mathbf{G}^{\tau-t}](a_{m+1}, a) \pi(a) [\mathbf{G}^{T-\tau} \mathbf{v}](a)}{\sum_b [\mathbf{G}^{\tau-t+1}](a_m, b) \pi(b) [\mathbf{G}^{T-\tau} \mathbf{v}](b)}$$

$$= p(a_m \rightarrow a_{m+1}; t) \quad \rightarrow \text{Markov process}$$

Time homogenous master equation

- ❖ Again, for infinite duration of the constraints, the transition probability $p(a \rightarrow b; t)$ for $\tau \geq t$ is independent of time, and independent of the initial distribution $\pi(a)$.
- ❖ However, the state occupation probability $p(a; t)$ is time-dependent, with initial condition $p(a; t = \tau) = \pi(a)$
- ❖ Discrete time-homogeneous master equation:

$$p(a; t + 1) = \sum_b p(b; t) p(b \rightarrow a)$$

Generalization



- ❖ Constraints on up to $(n+1)$ -point probability leads to n -th point Markov process
- ❖ Condition for time-homogeneity ← Generalization of Perron-Frobenius theorem to higher rank tensor required.

Summary



- ❖ Path entropy maximization: Most unbiased estimated of the path probability under the given constraint
- ❖ In particular, no correlations exists except those given by the constraints => n-th order Markov process if only up to n-point function is constrained.
- ❖ <http://arxiv.org/abs/1206.1416>
- ❖ Collaboration with Steve Pressé (UCSF)